

Statistical study of spatial dependences in
long memory random fields on a lattice,
point processes and random geometry.

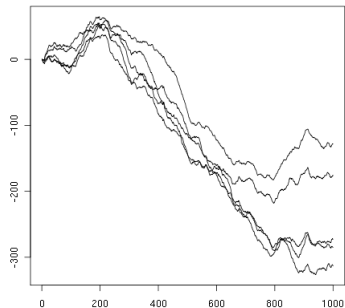
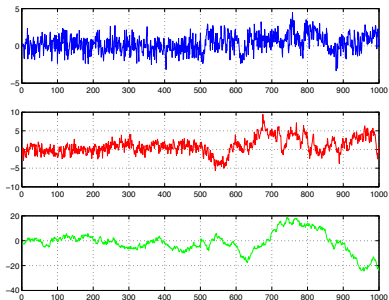
Frédéric Lavancier,

Laboratoire de Mathématiques Jean Leray, Nantes

9 décembre 2011.

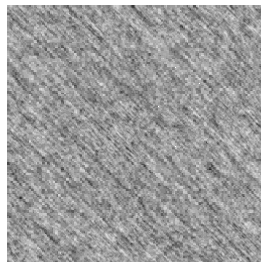
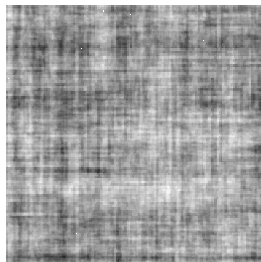
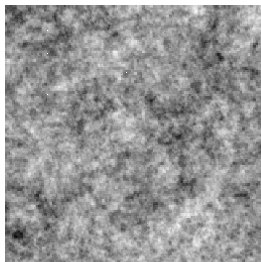
Introduction

Long memory, self-similarity in (multivariate) time series :



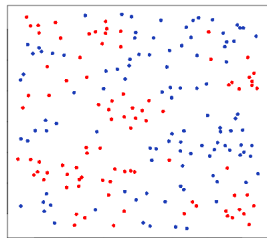
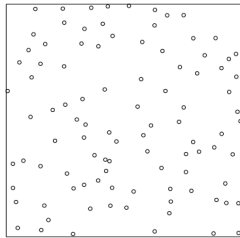
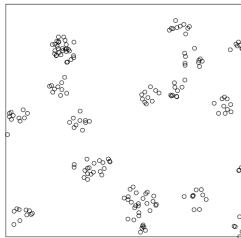
Introduction

Long memory, self-similarity in images :



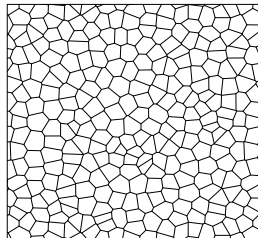
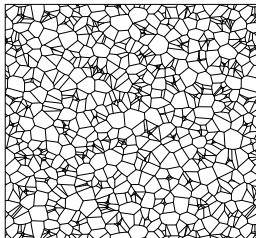
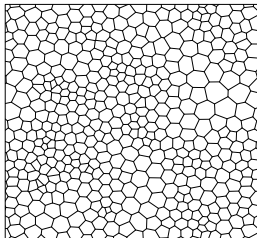
Introduction

Attraction, repulsion between points :



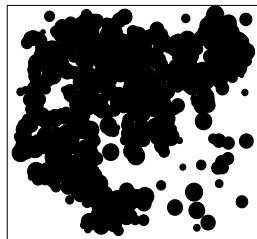
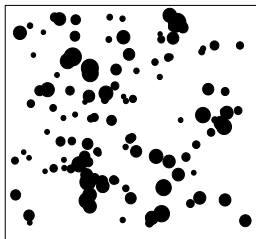
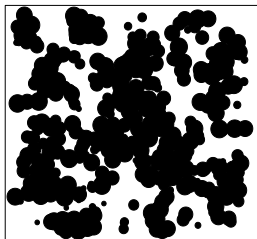
Introduction

Dependence between cells of a tessellation :



Introduction

Random sets as a union of interacting balls :



Introduction

General motivations

Developping mathematical models that

- respect some prescribed features (e.g. self-similarity, long range dependence, repulsion or attraction between points,...) ;
- are flexible enough (through few parameters) ;
- we can simulate.

Introduction

General motivations

Developping mathematical models that

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From a stastical point of view :

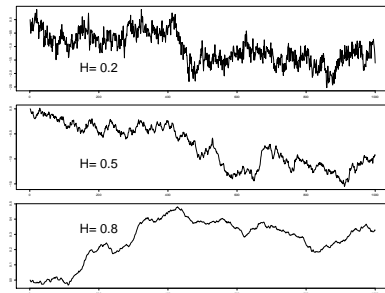
- fitting these models to data (inference problem);
- assessing the theoretical quality of inference (consistency, limiting law, optimality,...);
- providing some diagnostic tools :
 - adequation of the model to data,
 - change-point problems.

- 1 Self-similarity, Long memory
 - Vector Fractional Brownian Motion
 - Long memory time series
 - Long memory random fields (images)
- 2 Point processes, random geometry
 - Estimation of Gibbs point processes
 - Model validation for Gibbs point processes

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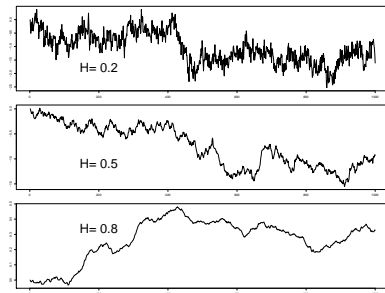
Fractional Brownian Motion

Some univariate examples :

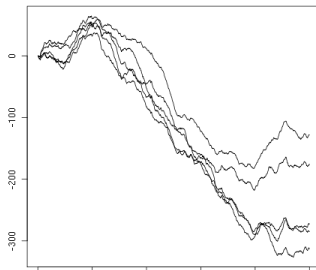


Fractional Brownian Motion

Some univariate examples :



Aim : correlating p fractional Brownian motions
 → Vector (or Multivariate) FBM



Vector FBM

with P.-O. Amblard, J.-F. Coeurjolly, A. Philippe, D. Surgailis

A multivariate process $\mathbf{B}(t) = (B_1(t), \dots, B_p(t))$ is a Vector FBM with parameter $\mathbf{H} = (H_1, \dots, H_p)$ if $\mathbf{B}(0) = \mathbf{0}$ and

- it is Gaussian;
- it is \mathbf{H} -self-similar, i.e.

$$\forall c > 0, (B_1(ct), \dots, B_p(ct))_{t \in \mathbb{R}} \stackrel{\mathcal{L}}{=} (c^{H_1} B_1(t), \dots, c^{H_p} B_p(t))_{t \in \mathbb{R}};$$

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Lamperti type result :

If there exist a vector process $(Y_1(t), \dots, Y_p(t))_{t \in \mathbb{R}}$ and real functions a_1, \dots, a_p such that

$$(a_1(n)Y_1(nt), \dots, a_p(n)Y_p(nt)) \underset{fidi}{\xrightarrow{n \rightarrow \infty}} \mathbf{Z}(t),$$

then the vector process $(\mathbf{Z}(t))$ is self-similar.

Convergence of partial sums :

Any Vector FBM can be obtained as the limit of partial sums of some superlinear processes.

Comprehensive characterisation : ($p = 2$)

1. Let $(B_1(t), B_2(t))$ be a (H_1, H_2) -VFBM, then (when $H_1 + H_2 \neq 1$) :

- B_1 is a H_1 -FBM and B_2 is a H_2 -FBM
- for $0 \leq s \leq t$, the cross-covariance is

$$\mathbb{E}B_1(s)B_2(t) \propto (\rho + \eta)s^{H_1+H_2} + (\rho - \eta)t^{H_1+H_2} - (\rho - \eta)(t - s)^{H_1+H_2}$$

$$\text{with } \rho = \text{corr}(B_1(1), B_2(1)), \quad \eta = \frac{\text{corr}(B_1(1), B_2(-1)) - \text{corr}(B_1(-1), B_2(1))}{2 - 2^{H_1+H_2}}.$$

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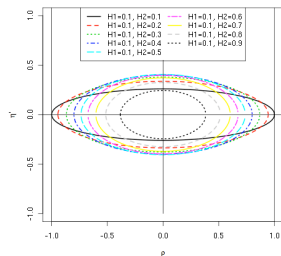
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2. Conversely any Gaussian process with the above covariance function is a VFBM iff for some known $R > 0$

$$\rho^2 \sin\left(\frac{\pi}{2}(H_1 + H_2)\right)^2 + \eta^2 \cos\left(\frac{\pi}{2}(H_1 + H_2)\right)^2 \leq R$$

→ not possible to set up arbitrary correlated FBM



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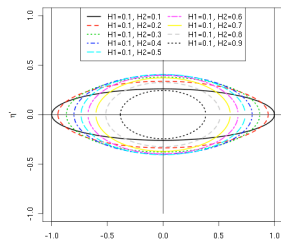
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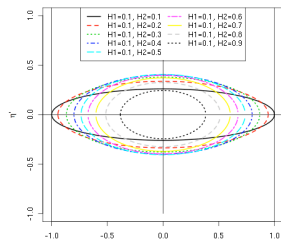
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The two components are either non-correlated components or

$$\mathbb{E}\Delta B_1(n)\Delta B_2(n+h) \sim \kappa|h|^{H_1+H_2-2}$$

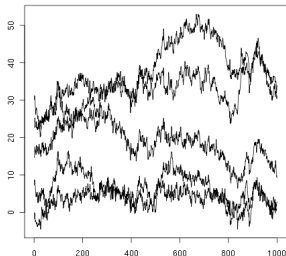
→ Very constrained cross-correlation

(For instance $H_1 + H_2 \geq 1 \Rightarrow$ long-range cross-dependence)

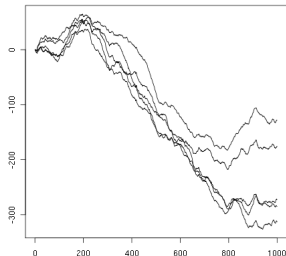
Some examples :

Simulations are achieved thanks to a Wood and Chan algorithm.

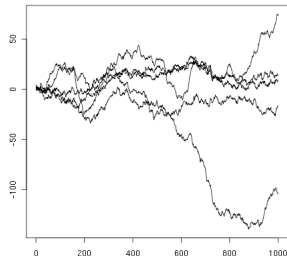
Below, the correlation between any couple of FBM is $\rho = 0.6$



$H \in [0.3, 0.4]$
(decentered)



$H \in [0.8, 0.9]$



$H \in [0.4, 0.8]$

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Definition :

A stationary, L^2 , time series $(X(n))_{n \in \mathbb{Z}}$ exhibits long memory if

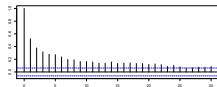
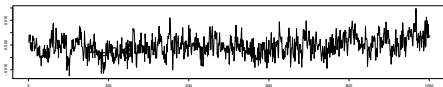
$$\sum_{n \in \mathbb{Z}} |r(n)| = +\infty$$

where r denotes the covariance function of X .

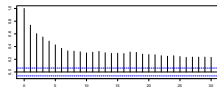
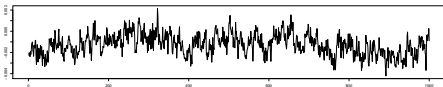
Typically : for some $0 < d < 0.5$, $r(n) \sim_{\infty} \kappa |n|^{2d-1}$.

The parameter d is called the *long memory parameter*.

$d = 0.3 :$



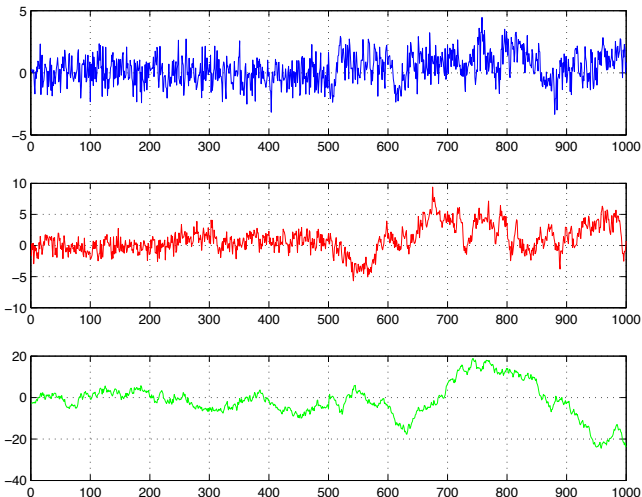
$d = 0.45 :$

**Examples :**

- If B is a FBM, $X(n) = B(n+1) - B(n)$ with $d = H - 0.5$.
- An $I(d)$ time series : $X(n) = (1 - L)^{-d} \epsilon(n)$
(where ϵ is a white noise and L the lag operator : $L\epsilon_n = \epsilon_{n-1}$)

Change point problem

with R. Leipus, A. Philippe, D. Surgailis

Constant vs non-constant long memory parameter? $\underline{ex} : I(d_1) \rightarrow I(d_2)$ 

Change point problem

with R. Leipus, A. Philippe, D. Surgailis

Basically, under \mathbf{H}_0 : d is constant :

$$n^{-d-0.5} \sum_{k=1}^{[nt]} X(k) \xrightarrow{\mathcal{D}([0,1])} \kappa B_{d+0.5}(t) \quad \text{and} \quad \text{Var} \left(\sum_{k=1}^{[nt]} X(k) \right) \approx n^{2d+1}$$

Under \mathbf{H}_1 : d increases, for some t_0

$$\text{Var} \left(\sum_{k=1}^{[nt_0]} X(k) \right) \ll \text{Var} \left(\sum_{k=[nt_0]}^n X(k) \right)$$

\Rightarrow To test for \mathbf{H}_0 against \mathbf{H}_1 , we estimate and compare these variances.

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Let $S_j = \sum_{k=1}^j X(k)$ and $S_{n-j}^* = \sum_{k=j+1}^n X(k)$. Define

- the *forward* variance : $V_k = \widehat{\text{Var}}(S_1, \dots, S_k)$
- the *backward* variance : $V_{n-k}^* = \widehat{\text{Var}}(S_{n-k+1}^*, \dots, S_1^*)$

Test statistic, Consistency

$$I_n = \int_0^1 \frac{V_{n-[nt]}^*}{V_{[nt]}} dt$$

Under \mathbf{H}_0 : $I_n \rightarrow I(B_{d+0.5})$ (simulable); Under \mathbf{H}_1 : $I_n \rightarrow +\infty$

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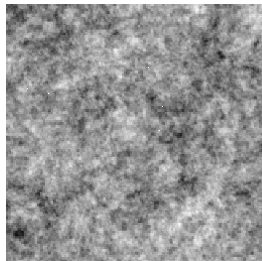
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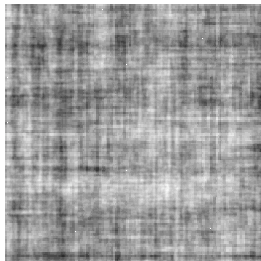
Main difference with time series : possible occurrence of anisotropy.

Examples ($d = 2$) : for $0 < \alpha < 1$

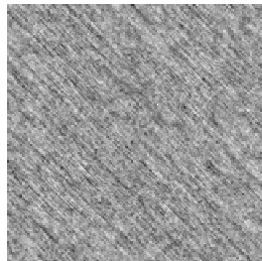
$$r(n_1, n_2) \sim_{\infty} (n_1^2 + n_2^2)^{-\alpha}$$



$$r(n_1, n_2) \sim_{\infty} |n_1|^{-\alpha} |n_2|^{-\alpha}$$



$$r(n_1, n_2) \sim_{\infty} |n_1 + n_2|^{-\alpha}$$



Investigations on long memory random fields

Modelling : long memory random fields appear

- in similar models as for time series (increment of fractional Brownian sheet ; aggregation of short memory random fields ; fractional filtering of a white noise)
- in some Gibbs processes on \mathbb{Z}^d in phase transition
ex : Ising model at the critical temperature

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Limit theorems, statistical applications :

- for partial sums \longrightarrow non Central Limit Theorems
 \Rightarrow Testing for the presence of long memory.
- for the empirical process \longrightarrow asymptotic degeneracy
 \Rightarrow Asymptotics for U-Statistics.
- for some quadratic forms (with A. Philippe)
 \longrightarrow non-CLTs for $\sum g(i-j)X(i)X(j)$ where $(X(i))_{i \in \mathbb{Z}^2}$ is Gaussian.
 \Rightarrow Asymptotics of empirical covariance functions.

Partial sums of long memory random fields

Let $X(n) = \sum_{k \in \mathbb{Z}^d} a_k \epsilon(n-k)$, where $(a_k) \in \ell^2$ and ϵ is a Gaussian white noise.

Theorem

Denote $a(x) = \sum_{k \in \mathbb{Z}^d} a_k e^{i k \cdot x}$. If $a \in L^2$ and $\forall \lambda, a(\lambda x) = |\lambda|^{-\alpha} a(x)$, $0 < \alpha < d$, then denoting $A_n = \{1, \dots, n\}^d$

$$\frac{1}{n^{d/2+\alpha}} \sum_{k \in A_{[nt]}} X_k \xrightarrow{\mathcal{D}([0,1]^d)} \int_{\mathbb{R}^d} a(x) \prod_{j=1}^d \frac{e^{it_j x_j} - 1}{ix_j} dZ(x)$$

The limit is the Fractional Brownian Sheet only when $a(x) = \prod_{i=1}^d |x_i|^{-H_i}$.

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$$\begin{aligned} \frac{1}{n^{d/2+\alpha}} \sum_{k \in A_n} X_k &= \frac{1}{n^{d/2+\alpha}} \sum_{k \in A_n} \int_{[-\pi, \pi]^d} a(x) e^{i k \cdot x} dZ(x) \quad Z : \text{spectral measure of } \epsilon \\ &= \int_{[-\pi, \pi]^d} n^{-\alpha} a(x) \sum_{k \in A_n} e^{i k \cdot x} \frac{dZ(x)}{n^{d/2}} \\ &= \int_{[-n\pi, n\pi]^d} a(x) \prod_{j=1}^d \frac{e^{ix_j} - 1}{n(e^{ix_j/n} - 1)} dZ(x) \xrightarrow{L^2} \int_{\mathbb{R}^d} a(x) \prod_{j=1}^d \frac{e^{ix_j} - 1}{ix_j} dZ(x) \end{aligned}$$

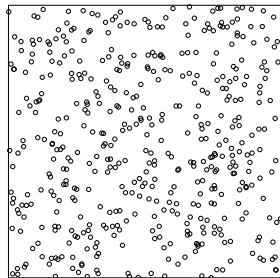
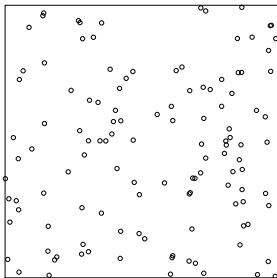
- ② Point processes, random geometry
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Point processes

Notation :

- Denote by φ a point pattern on \mathbb{R}^d , i.e. $\varphi = \bigcup_{i \in \mathcal{I}} x_i$, for $\mathcal{I} \subset \mathbb{N}^*$
- A point process Φ is a random variable on the space $\Omega = \{\varphi\}$.

Example : the Poisson point process \rightarrow independance in locations



How to introduce dependencies between the location of points ?

Gibbs point processes (basic definition)

They are absolutely continuous w.r.t the Poisson point process with density

$$f_{\theta}(\varphi) = \frac{1}{c_{\theta}} e^{-V_{\theta}(\varphi)}, \quad \text{for some parameter } \theta \in \mathbb{R}^p.$$

$V_{\theta}(\varphi)$: energy of φ also called Hamiltonian (belongs to $\mathbb{R} \cup \{+\infty\}$)

- φ is more likely to occur if $V_{\theta}(\varphi)$ is small.
- if $V_{\theta}(\varphi) = +\infty$, then φ is forbidden.

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They are absolutely continuous w.r.t the Poisson point process with density

$$f_{\theta}(\varphi) = \frac{1}{c_{\theta}} e^{-V_{\theta}(\varphi)}, \quad \text{for some parameter } \theta \in \mathbb{R}^p.$$

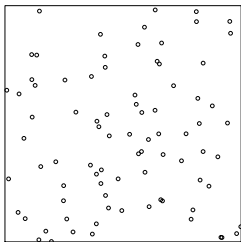
$V_{\theta}(\varphi)$: energy of φ also called Hamiltonian (belongs to $\mathbb{R} \cup \{+\infty\}$)

- φ is more likely to occur if $V_{\theta}(\varphi)$ is small.
- if $V_{\theta}(\varphi) = +\infty$, then φ is forbidden.

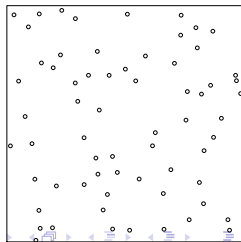
Example : Strauss process with range of interaction $R = 0.05$ on $[0, 1]^2$

$$V_{\theta}(\varphi) = \theta \sum_{(x,y) \in \varphi} \mathbb{I}_{|y-x| < R}, \quad \theta > 0$$

$\theta = 0.35$:



$\theta = 2.3$:



Geometric interactions

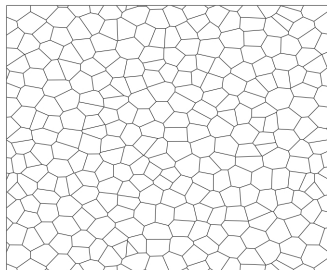
Geometric interactions

For a point pattern φ , denote $\text{Vor}(\varphi)$ the associated Voronoï tessellation.

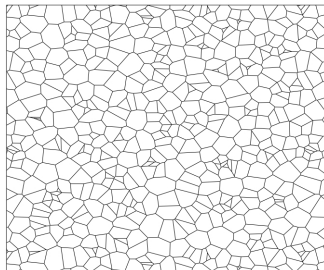
Gibbs Voronoï tessellation : one example

$$V_\theta(\varphi) = \sum_{C \in \text{Vor}(\varphi)} V_1(C) + \theta \sum_{\substack{C, C' \in \text{Vor}(\varphi) \\ C \text{ and } C' \text{ are neighbors}}} |vol(C) - vol(C')|$$

$$V_1(C) = \begin{cases} +\infty & \text{if the cell is too "irregular",} \\ 0 & \text{otherwise} \end{cases}$$



$\theta > 0$



$\theta < 0$

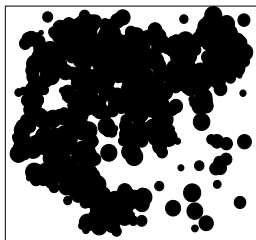
Geometric interactions

Quermass model

Each point $x \in \varphi$ is associated with a random radius r .

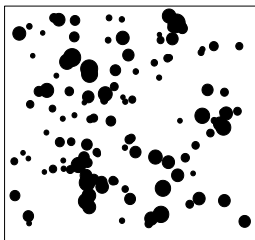
$$V_\theta(\varphi) = \theta_1 \mathcal{P}(\Gamma) + \theta_2 \mathcal{A}(\Gamma) + \theta_3 \mathcal{E}(\Gamma) \quad \text{where} \quad \Gamma = \bigcup_{x \in \varphi} B(x, r)$$

\mathcal{P} : perimeter \mathcal{A} : area \mathcal{E} : EP characteristic (nb connected sets - nb holes).



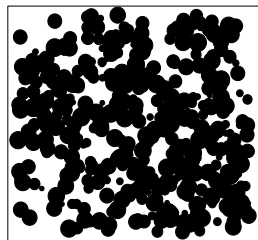
$$\theta_1 > 0$$

$$(\theta_2 = \theta_3 = 0)$$



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Statistical issues

Let φ be a point pattern observed (or $\text{Vor}(\varphi)$, or Γ, \dots) on a domain Λ .

Assumption : φ is the realisation of a Gibbs point process associated to V_θ for some (unknown) $\theta \in \mathbb{R}^p$.

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❷ Is the above assumption reasonable?

Takacs-Fiksel method

Campbell equilibrium equation (Georgii, Nguyen, Zessin)

Let $V_\theta(x|\varphi) = V_\theta(\varphi \cup x) - V_\theta(\varphi)$ (energy needed to insert x in φ)

Φ is a Gibbs point process with energy V_θ if and only if, for all function h ,

$$\mathbb{E}_\Phi \left(\int_{\mathbb{R}^d} h(x, \Phi) e^{-V_\theta(x|\Phi)} dx \right) = \mathbb{E}_\Phi \left(\sum_{x \in \Phi} h(x, \Phi \setminus x) \right).$$

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Empirical counterpart : **Takacs Fiksel estimation**

Let h_1, \dots, h_k be K test functions (to be chosen),

$$\hat{\theta} = \arg \min_{\theta} \sum_{k=1}^K \left[\int_{\Lambda} h_k(x, \varphi) e^{-V_\theta(x|\varphi)} dx - \sum_{x \in \varphi} h_k(x, \varphi \setminus x) \right]^2.$$

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TF estimation includes pseudo-likelihood estimation as a particular case

Contributions : with J.-F. Coeurjolly, D. Dereudre, R. Drouihet, K. Stankova-Helisova

- Identifiability ($K > \dim(\theta)$), consistency, asymptotic normality ;
- Extension to a two-step procedure in presence of (possible non-hereditary) hardcore interactions ;
- Application to Gibbs Voronoï tessellations ;
- Application to Quermass model, where points are not observed.

Model Validation

with J.-F. Coeurjolly

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The residuals assess the Campbell equilibrium : for any h ,

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If the model is well specified, we expect, for any h , $R_\Lambda(h) \approx 0$.

We have proved : as $\Lambda \rightarrow \mathbb{R}^d$, $R_\Lambda(h) \rightarrow 0$ and $R_\Lambda(h) \sim \mathcal{N}(0, \Sigma)$.

Towards χ^2 goodness of fit tests : 2 frameworks

$$\begin{cases} R_\Lambda(h_1) \\ \vdots \\ R_\Lambda(h_j) \\ \vdots \\ R_\Lambda(h_s) \end{cases}$$

$$\sum_{j=1}^s R_\Lambda(\varphi, h_j)^2 \sim \chi^2$$

$R_{\Lambda_1}(h)$	$R_{\Lambda_2}(h)$	\dots
\dots	\dots	\dots
\dots	\dots	$R_{\Lambda_q}(h)$

$$\sum_{j=1}^q R_{\Lambda_j}(\varphi, h)^2 \sim \chi^2$$

Theoretical Ingredients

For the estimation and the validation procedures, our asymptotic results rely on

- the Campbell equilibrium equation,
- an ergodic theorem (Nguyen and Zessin),
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- the central limit theorem below.

CLT for a linear functional $U(\Phi_\Lambda)$ where $\Phi_\Lambda = \Phi \cap \Lambda$

Assume that if $\Lambda = \cup_{k=1}^n \Delta_k$ for disjoint Δ_k 's then $U(\Phi_\Lambda) = \sum_{k=1}^n U(\Phi_{\Delta_k})$

Basically, if

(i) Conditioned centering : $\mathbb{E} [U(\Phi_{\Delta_k}) | \Phi_{\Delta_j}, j \neq k] = 0$

(ii) Convergence of empirical covariances : $\frac{1}{n} \sum_{k=1}^n \sum_{k'=1}^n U(\Phi_{\Delta_k}) U(\Phi_{\Delta_{k'}}) \xrightarrow{L^1} \Sigma$

then $\frac{1}{\sqrt{n}} \sum_{k=1}^n U(\Phi_{\Delta_k}) \xrightarrow{\mathcal{L}} \mathcal{N}(0, \Sigma)$

The key assumption (i) allows us to go **without mixing assumptions** (which typically do not hold for all values of θ for a Gibbs process).

Alternatives to Gibbs point processes?

Gibbs processes introduce interactions in a very natural way, but

- they can be tedious to simulate
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Alternatives to Gibbs processes :

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Alternatives to Gibbs processes :

- **Cox processes** : Poisson processes with random intensity function.
⇒ They induce clustered point patterns.
- **Determinantal point processes** (with J. Møller and E. Rubak)
Their joint intensities depend on a covariance function C .
⇒ They are repulsive point processes, for $g(h) = 1 - \frac{C(h)^2}{C(0)^2}$.

Appeals :

- ▷ Perfect and fast simulation is available
- ▷ Flexible models : just consider a parametric family of covariance functions
- ▷ Inference is feasible by standard methods (maximum likelihood, contrast functions,...)

Thank you.